## General Aptitude (GA)

## Q. 1 - Q. 5 Carry ONE mark Each

| Q. 1 | "I have not yet decided what I will do this evening; I____ visit a friend." |
| :--- | :--- |
|  |  |
| (A) | mite |
| (B) | would |
| (C) | might |
| (D) | didn't |
|  |  |


| Q.2 | Eject : Insert : : Advance : _____ <br> (By word meaning) |
| :--- | :--- |
|  |  |
| (A) | Advent |
| (B) | Progress |
| (C) | Retreat |
| (D) | Loan |
|  |  |


| Q.3 | In the given figure, PQRSTV is a regular hexagon with each side of length 5 cm . A <br> circle is drawn with its centre at V such that it passes through P. What is the area <br> (in $\mathrm{cm}^{2}$ ) of the shaded region? (The diagram is representative) |
| :--- | :--- |
|  |  |
| (A) | $\frac{25 \pi}{3}$ |
| (B) | $\frac{20 \pi}{3}$ |
| (C) | $6 \pi$ |
| (D) | $7 \pi$ |


| Q.4 | A duck named Donald Duck says "All ducks always lie." <br> Based only on the information above, which one of the following statements can be <br> logically inferred with certainty? |
| :--- | :--- |
| (A) | Donald Duck always lies. |
| (B) | Donald Duck always tells the truth. |
| (C) | Donald Duck's statement is true. |
| (D) | Donald Duck's statement is false. |
|  |  |


| Q. 5 | A line of symmetry is defined as a line that divides a figure into two parts in a way <br> such that each part is a mirror image of the other part about that line. <br> The figure below consists of 20 unit squares arranged as shown. In addition to the <br> given black squares, upto 5 more may be coloured black. Which one among the <br> following options depicts the minimum number of boxes that must be coloured <br> black to achieve two lines of symmetry? (The figure is representative) |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |

## Q. 6 - Q. 10 Carry TWO marks Each

| Q.6 | Based only on the truth of the statement 'Some humans are intelligent', which one <br> of the following options can be logically inferred with certainty? |
| :--- | :--- |
|  |  |
| (A) | No human is intelligent. |
| (B) | All humans are intelligent. |
| (C) | Some non-humans are intelligent. |
| (D) | Some intelligent beings are humans. |


| Q.7 | Which one of the options can be inferred about the mean, median, and mode for the <br> given probability distribution (i.e. probability mass function), $P(x)$, of a variable $x$ ? |
| :--- | :--- |
|  |  |
| (A) | mean $=$ median $\neq$ mode |
| (B) | mean $=$ median $=$ mode |
| (C) | mean $\neq$ median $=$ mode |
| (D) | mean $\neq$ mode $=$ median |


| Q. 8 | The James Webb telescope, recently launched in space, is giving humankind <br> unprecedented access to the depths of time by imaging very old stars formed almost <br> 13 billion years ago. Astrophysicists and cosmologists believe that this odyssey in <br> space may even shed light on the existence of dark matter. Dark matter is supposed <br> to interact only via the gravitational interaction and not through the <br> electromagnetic-, the weak- or the strong-interaction. This may justify the epithet <br> "dark" in dark matter. <br> Based on the above paragraph, which one of the following statements is FALSE? |
| :--- | :--- |
| (A) | No other telescope has captured images of stars older than those captured by the <br> James Webb telescope. |
| (B) | People other than astrophysicists and cosmologists may also believe in the existence <br> of dark matter. |
| (C) | The James Webb telescope could be of use in the research on dark matter. |
| (D) | If dark matter was known to interact via the strong-interaction, then the epithet <br> "dark" would be justified. |
|  |  |


| Q. 9 | Let $a=30!, b=50!$, and $c=100!$. Consider the following numbers: $\log _{a} c, \quad \log _{c} a, \quad \log _{b} a, \quad \log _{a} b$ <br> Which one of the following inequalities is CORRECT? |
| :---: | :---: |
|  |  |
| (A) | $\log _{c} a<\log _{b} a<\log _{a} b<\log _{a} c$ |
| (B) | $\log _{c} a<\log _{a} b<\log _{b} a<\log _{b} c$ |
| (C) | $\log _{c} a<\log _{b} a<\log _{a} c<\log _{a} b$ |
| (D) | $\log _{b} a<\log _{c} a<\log _{a} b<\log _{a} c$ |
|  |  |


| Q. 10 | A square of side length 4 cm is given. The boundary of the shaded region is defined <br> by one semi-circle on the top and two circular arcs at the bottom, each of radius <br> 2 cm, as shown. <br> The area of the shaded region is __ |
| :--- | :--- |
| (A) | 8 |
| (C) |  |

## Q. 11 - Q. 35 Carry ONE mark Each

| Q.11 | The area of the region bounded by the parabola $x=-y^{2}$ and the line <br> $y=x+2$ equals |
| :--- | :--- |
| (A) | $\frac{3}{2}$ |
| (B) | $\frac{7}{2}$ |
| (C) | $\frac{9}{2}$ |
| (D) | 9 |
| Q.12 | Let $A$ be a $3 \times 3$ real matrix having eigenvalues 1,0, and -1. <br> If $B=A^{2}+2 A+I_{3}$, where $I_{3}$ is the $3 \times 3$ identity matrix, then which one <br> of the following statements is true? |
| (B) | $B^{3}-5 B^{2}-4 B=0$ |
| (C) | $B^{3}+5 B^{2}-4 B=0$ |
|  | $B^{3}-5 B^{2}+4 B=0$ |
|  |  |
| (A) |  |


| Q.13 | Consider the following statements. <br> (I) $\quad$Let $A$ and $B$ be two $n \times n$ real matrices. If $B$ is invertible, then <br> $\operatorname{rank}(B A)=\operatorname{rank}(A)$. <br> Let $A$ be an $n \times n$ real matrix. If $A^{2} \boldsymbol{x}=\boldsymbol{b}$ has a solution for every <br> $\boldsymbol{b} \in \mathbb{R}^{n}$, then $A \boldsymbol{x}=\boldsymbol{b}$ also has a solution for every $\boldsymbol{b} \in \mathbb{R}^{n}$. <br> Which of the above statements is/are true? |
| :--- | :--- |
| (A) | Only (I) <br> (B) <br> Only (II) <br> (D) and (II) |


| Q. 14 | Consider the probability space $(\Omega, \mathcal{G}, P)$, where $\Omega=[0,2]$ and $\mathcal{G}=\{\phi, \Omega,[0,1],(1,2]\}$. Let $X$ and $Y$ be two functions on $\Omega$ defined as $X(\omega)= \begin{cases}1 & \text { if } \omega \in[0,1] \\ 2 & \text { if } \omega \in(1,2]\end{cases}$ <br> and $Y(\omega)= \begin{cases}2 & \text { if } \omega \in[0,1.5] \\ 3 & \text { if } \omega \in(1.5,2]\end{cases}$ <br> Then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | $X$ is a random variable with respect to $\mathcal{G}$, but $Y$ is not a random variable with respect to $\mathcal{G}$ |
| (B) | $Y$ is a random variable with respect to $\mathcal{G}$, but $X$ is not a random variable with respect to $\mathcal{G}$ |
| (C) | Neither $X$ nor $Y$ is a random variable with respect to $\mathcal{G}$ |
| (D) | Both $X$ and $Y$ are random variables with respect to $\mathcal{G}$ |
|  |  |


| Q.15 | Let $\Phi(\cdot)$ denote the cumulative distribution function of a standard normal <br> random variable. If the random variable $X$ has the cumulative distribution <br> function |
| :--- | :--- |
| $\qquad F(x)= \begin{cases}\Phi(x) & \text { if } x<-1 \\ \Phi(x+1) & \text { if } x \geq-1,\end{cases}$ |  |
| then which one of the following statements is true? |  |$\quad$| (A) | $P(X \leq-1)=\frac{1}{2}$ |
| :--- | :--- |
| (B) | $P(X=-1)=\frac{1}{2}$ |
| (D) | $P(X \leq 0)=\frac{1}{2}$ |
|  |  |


| Q. 16 | Let $X$ be a random variable with probability density function $f(x)= \begin{cases}\alpha \lambda x^{\alpha-1} e^{-\lambda x^{\alpha}} & \text { if } x>0 \\ 0 & \text { otherwise }\end{cases}$ <br> where $\alpha>0$ and $\lambda>0$. If the median of $X$ is 1 and the third quantile is 2 , then $(\alpha, \lambda)$ equals |
| :---: | :---: |
|  |  |
| (A) | $\left(1, \log _{e} 2\right)$ |
| (B) | $(1,1)$ |
| (C) | $\left(2, \log _{e} 2\right)$ |
| (D) | $\left(1, \log _{e} 3\right)$ |
|  |  |


| Q. 17 | Let $X$ be a random variable having Poisson distribution with mean $\lambda>0$. Then <br> $E\left(\left.\frac{1}{X+1} \right\rvert\, X>0\right)$ equals |
| :--- | :--- |
| (A) | $\frac{1-e^{-\lambda}-\lambda e^{-\lambda}}{\lambda\left(1-e^{-\lambda}\right)}$ |
| (B) | $\frac{1-e^{-\lambda}}{\lambda}$ |
| (C) | $\frac{1-e^{-\lambda}-\lambda e^{-\lambda}}{\lambda}$ |
| (D) | $\frac{1-e^{-\lambda}}{\lambda+1}$ |


| Q.18 | Suppose that $X$ has the probability density function |
| :--- | :--- |
| $\qquad$ | $f(x)= \begin{cases}\frac{\lambda^{\alpha}}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} \quad \text { if } x>0 \\ 0 & \text { where } \alpha>0 \text { and } \lambda>0 . \text { Which one of the following statements is NOT true? }\end{cases}$ |
| (A) | $E(X)$ exists for all $\alpha>0$ and $\lambda>0$ |
| (B) | Variance of $X$ exists for all $\alpha>0$ and $\lambda>0$ |
| (C) | $E\left(\frac{1}{X}\right)$ exists for all $\alpha>0$ and $\lambda>0$ |
| (D) | $E\left(\log _{e}(1+X)\right)$ exists for all $\alpha>0$ and $\lambda>0$ |
|  |  |

\(\left.$$
\begin{array}{|l|l|}\hline \text { Q.19 } & \text { Let }(X, Y) \text { have joint probability density function } \\
\qquad & f(x, y)=\left\{\begin{array}{ll|}8 x y & \text { if } 0<x<y<1 \\
0 & \text { otherwise. }\end{array}
$$\right. <br>

\hline (A) E\left(X \mid Y=y_{0}\right)=\frac{1}{2}, then y_{0} equals\end{array}\right\}\)| (B) | $\frac{3}{4}$ |
| :--- | :--- |
| (C) | $\frac{1}{2}$ |
| (D) | $\frac{2}{3}$ |
|  |  |

\(\left.$$
\begin{array}{|l|l|}\hline \text { Q.20 } & \begin{array}{l}\text { Suppose that there are } 5 \text { boxes, each containing } 3 \text { blue pens, } 1 \text { red pen and } 2 \\
\text { black pens. One pen is drawn at random from each of these } 5 \text { boxes. If the random } \\
\text { variable } X_{1} \text { denotes the total number of blue pens drawn and the random variable } \\
X_{2} \text { denotes the total number of red pens drawn, then } P\left(X_{1}=2, X_{2}=1\right) \text { equals }\end{array} \\
\hline \text { (A) } & \frac{5}{36} \\
\hline \text { (B) } & \frac{5}{18} \\
\hline \text { (C) } & \frac{5}{12} \\
\hline \text { (D) } & \begin{array}{l}\frac{5}{9} \\
\hline \text { Q.21 }\end{array}
$$ <br>
\hline Let\left\{X_{n}\right\}_{n \geq 1} and\left\{Y_{n}\right\}_{n \geq 1} be two sequences of random variables and X and Y <br>
be two random variables, all of them defined on the same probability space. <br>

Which one of the following statements is true?\end{array}\right\}\)| If $\left\{E\left(X_{n}\right)\right\}_{n \geq 1}$ converges to $E(X)$, then $\left\{X_{n}\right\}_{n \geq 1}$ converges in $1^{\text {st mean to } X}$ |
| :--- |
| (A) |
| If $\left\{X_{n}\right\}_{n \geq 1}$ converges in distribution to $X$ and $\left\{Y_{n}\right\}_{n \geq 1}$ converges in |
| (D $\left\{X_{n}\right\}_{n \geq 1}$ converges in distribution to a real constant $c$, then $\left\{X_{n}\right\}_{n \geq 1}$ |
| converges in probability to $c$ |


| Q.22 | Let $X$ be a random variable with probability density function |
| :--- | :--- |
| $\qquad$ $f(x ; \lambda)= \begin{cases}\frac{1}{\lambda} e^{-\frac{x}{\lambda}} & \text { if } x>0 \\ 0 & \text { otherwise, } \\ \text { where } \lambda>0 \text { is an unknown parameter. Let } Y_{1}, Y_{2}, \ldots, Y_{n} \text { be a random sample of } \\ \text { size } n \text { from a population having the same distribution as } X^{2} . \\ \text { If } \bar{Y}=\frac{1}{n} \sum_{i=1}^{n} Y_{i}, \text { then which one of the following statements is true? }\end{cases}$ <br> (A) $\sqrt{\frac{\bar{Y}}{2}}$ is a method of moments estimator of $\lambda$ |  |
| (B) | $\sqrt{\bar{Y}}$ is a method of moments estimator of $\lambda$ |
| (C) | $\frac{1}{2} \sqrt{\bar{Y}}$ is a method of moments estimator of $\lambda$ |
| (D) | $2 \sqrt{\bar{Y}}$ is a method of moments estimator of $\lambda$ |
|  |  |


| Q. 23 | Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample of size $n(\geq 2)$ from a population having probability density function $f(x ; \theta)= \begin{cases}\frac{2}{\theta x}\left(-\log _{e} x\right) e^{-\frac{\left(\log _{e} x\right)^{2}}{\theta}} & \text { if } 0<x<1 \\ 0 & \text { otherwise }\end{cases}$ <br> where $\theta>0$ is an unknown parameter. Then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | $\frac{1}{n} \sum_{i=1}^{n}\left(\log _{e} X_{i}\right)^{2}$ is the maximum likelihood estimator of $\theta$ |
| (B) | $\frac{1}{n-1} \sum_{i=1}^{n}\left(\log _{e} X_{i}\right)^{2}$ is the maximum likelihood estimator of $\theta$ |
| (C) | $\frac{1}{n} \sum_{i=1}^{n} \log _{e} X_{i}$ is the maximum likelihood estimator of $\theta$ |
| (D) | $\frac{1}{n-1} \sum_{i=1}^{n} \log _{e} X_{i}$ is the maximum likelihood estimator of $\theta$ |
|  |  |


| Q.24 | Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample of size $n$ from a population having <br> uniform distribution over the interval $\left(\frac{1}{3}, \theta\right)$, where $\theta>\frac{1}{3}$ is an unknown <br> parameter. If $Y=\max \left\{X_{1}, X_{2}, \ldots, X_{n}\right\}$, then which one of the following <br> statements is true? |
| :--- | :--- |
| (A) | $\left(\frac{n+1}{n}\right)\left(Y-\frac{1}{3}\right)+\frac{1}{3}$ is an unbiased estimator of $\theta$ |
| (B) | $\left(\frac{n}{n+1}\right)\left(Y-\frac{1}{3}\right)+\frac{1}{3}$ is an unbiased estimator of $\theta$ |
| (C) | $\left(\frac{n+1}{n}\right)\left(Y+\frac{1}{3}\right)-\frac{1}{3}$ is an unbiased estimator of $\theta$ |
| (D) | $Y$ is an unbiased estimator of $\theta$ |
|  |  |


| Q.25 | Suppose that $\boldsymbol{X}_{1}, \boldsymbol{X}_{2}, \ldots, \boldsymbol{X}_{n}, \boldsymbol{Y}_{1}, \boldsymbol{Y}_{2}, \ldots, \boldsymbol{Y}_{n}$ are independent and identically <br> distributed random vectors each having $N_{p}(\boldsymbol{\mu}, \Sigma)$ distribution, where $\Sigma$ is non- <br> singular, $p>1$ and $n>1$. If $\overline{\boldsymbol{X}}=\frac{1}{n} \sum_{i=1}^{n} \boldsymbol{X}_{i}$ and $\overline{\boldsymbol{Y}}=\frac{1}{n} \sum_{i=1}^{n} \boldsymbol{Y}_{i}$, then which <br> one of the following statements is true? |
| :--- | :--- |
| (A) | There exists $c>0$ such that $c(\overline{\boldsymbol{X}}-\boldsymbol{\mu})^{T} \Sigma^{-1}(\overline{\boldsymbol{X}}-\boldsymbol{\mu})$ has $\chi^{2}$-distribution <br> with $p$ degrees of freedom |
| (B) | There exists $c>0$ such that $c(\overline{\boldsymbol{X}}-\overline{\boldsymbol{Y}})^{T} \Sigma^{-1}(\overline{\boldsymbol{X}}-\overline{\boldsymbol{Y}})$ has $\chi^{2}$-distribution <br> with $(p-1)$ degrees of freedom |
| (C)There exists $c>0$ such that $c \sum_{i=1}^{n}\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}\right)^{T} \Sigma^{-1}\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}\right)$ has <br> $\chi^{2}$-distribution with $p$ degrees of freedom |  |
| (D) | There exists $c>0$ such that $c \sum_{i=1}^{n}\left(\boldsymbol{X}_{i}-\boldsymbol{Y}_{i}-\overline{\boldsymbol{X}}+\overline{\boldsymbol{Y}}\right)^{T} \Sigma^{-1}\left(\boldsymbol{X}_{i}-\boldsymbol{Y}_{i}-\overline{\boldsymbol{X}}+\overline{\boldsymbol{Y}}\right)$ <br> has $\chi^{2}$-distribution with $p$ degrees of freedom |
|  | (D) |


| Q. 26 | Consider the following regression model $y_{k}=\alpha_{0}+\alpha_{1} \log _{e} k+\epsilon_{k}, \quad k=1,2, \ldots, n$ <br> where $\epsilon_{k}$ 's are independent and identically distributed random variables each having probability density function $f(x)=\frac{1}{2} e^{-\|x\|}, x \in \mathbb{R}$. Then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | The maximum likelihood estimator of $\alpha_{0}$ does not exist |
| (B) | The maximum likelihood estimator of $\alpha_{1}$ does not exist |
| (C) | The least squares estimator of $\alpha_{0}$ exists and is unique |
| (D) | The least squares estimator of $\alpha_{1}$ exists, but it is not unique |
|  |  |


| Q. 27 | Suppose that $X_{1}, X_{2}, \ldots, X_{n}$ are independent and identically distributed random variables each having probability density function $f(\cdot)$ and median $\theta$. We want to test $H_{0}: \theta=\theta_{0} \quad \text { against } \quad H_{1}: \theta>\theta_{0}$ <br> Consider a test that rejects $H_{0}$ if $S>c$ for some $c$ depending on the size of the test, where $S$ is the cardinality of the set $\left\{i: X_{i}>\theta_{0}, 1 \leq i \leq n\right\}$. Then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | Under $H_{0}$, the distribution of $S$ depends on $f(\cdot)$ |
| (B) | Under $H_{1}$, the distribution of $S$ does not depend on $f(\cdot)$ |
| (C) | The power function depends on $\theta$ |
| (D) | The power function does not depend on $\theta$ |
|  |  |



| Q. 32 | Let $\left\{X_{n}\right\}_{n \geq 1}$ be a sequence of independent and identically distributed random variables each having mean 4 and variance 9. If $Y_{n}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$ for $n \geq 1$, then $\lim _{n \rightarrow \infty} E\left[\left(\frac{Y_{n}-4}{\sqrt{n}}\right)^{2}\right]$ (in integer) equals $\qquad$ |
| :---: | :---: |
| Q. 33 | Let $\left\{W_{t}\right\}_{t \geq 0}$ be a standard Brownian motion. Then $E\left(W_{4}^{2} \mid W_{2}=2\right)$ (in integer) equals $\qquad$ |
| Q. 34 | Let $\left\{X_{n}\right\}_{n \geq 1}$ be a Markov chain with state space $\{1,2,3\}$ and transition probability matrix $\left[\begin{array}{lll} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{array}\right] .$ <br> Then $P\left(X_{2}=1 \mid X_{1}=1, X_{3}=2\right)$ (rounded off to two decimal places) equals $\qquad$ |
| Q. 35 | Suppose that $\left(X_{1}, X_{2}, X_{3}\right)$ has $N_{3}(\boldsymbol{\mu}, \Sigma)$ distribution with $\boldsymbol{\mu}=\left[\begin{array}{l}0 \\ 0 \\ 0\end{array}\right]$ and $\Sigma=\left[\begin{array}{lll} 2 & 2 & 1 \\ 2 & 5 & 1 \\ 1 & 1 & 1 \end{array}\right]$ <br> Given that $\Phi(-0.5)=0.3085$, where $\Phi(\cdot)$ denotes the cumulative distribution function of a standard normal random variable, $P\left(\left(X_{1}-2 X_{2}+2 X_{3}\right)^{2}<\frac{7}{2}\right)$ (rounded off to two decimal places) equals $\qquad$ |

## Q. 36 - Q. 65 Carry TWO marks Each

| Q. 36 | Let $A$ be an $n \times n$ real matrix. Consider the following statements. <br> (I) If $A$ is symmetric, then there exists $c \geq 0$ such that $A+c I_{n}$ is symmetric and positive definite, where $I_{n}$ is the $n \times n$ identity matrix. <br> (II) If $A$ is symmetric and positive definite, then there exists a symmetric and positive definite matrix $B$ such that $A=B^{2}$. <br> Which of the above statements is/are true? |
| :---: | :---: |
|  |  |
| (A) | Only (I) |
| (B) | Only (II) |
| (C) | Both (I) and (II) |
| (D) | Neither (I) nor (II) |
|  |  |


| Q.37 | Let $X$ be a random variable with probability density function |
| :--- | :--- |
| $\qquad$  <br>  $f(x)= \begin{cases}\frac{1}{x^{2}} & \text { if } x \geq 1 \\ 0 & \text { otherwise. }\end{cases}$ <br> (A) $Y=\log _{e} X$, then $P(Y<1 \mid Y<2)$ equals  |  |
| (B) | $\frac{e-1}{1+e}$ |
| (C) | $\frac{1}{1+e}$ |
| (D) | $\frac{1}{e-1}$ |
|  |  |


| Q. 38 | Let $\{N(t)\}_{t \geq 0}$ be a Poisson process with rate 1 . Consider the following statements. <br> (I) $\quad P(N(3)=3 \mid N(5)=5)=\binom{5}{3}\left(\frac{3}{5}\right)^{3}\left(\frac{2}{5}\right)^{2}$. <br> (II) If $S_{5}$ denotes the time of occurrence of the $5^{\text {th }}$ event for the above Poisson process, then $E\left(S_{5} \mid N(5)=3\right)=7$. <br> Which of the above statements is/are true? |
| :---: | :---: |
|  |  |
| (A) | Only (I) |
| (B) | Only (II) |
| (C) | Both (I) and (II) |
| (D) | Neither (I) nor (II) |
|  |  |


| Q. 39 | Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample of size $n$ from a population having probability density function $f(x ; \mu)= \begin{cases}e^{-(x-\mu)} & \text { if } \mu \leq x<\infty \\ 0 & \text { otherwise }\end{cases}$ <br> where $\mu \in \mathbb{R}$ is an unknown parameter. If $\widehat{M}$ is the maximum likelihood estimator of the median of $X_{1}$, then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | $P(\widehat{M} \leq 2)=1-e^{-n\left(1-\log _{e} 2\right)}$ if $\mu=1$ |
| (B) | $P(\widehat{M} \leq 1)=1-e^{-n \log _{e} 2}$ if $\mu=1$ |
| (C) | $P(\widehat{M} \leq 3)=1-e^{-n\left(1-\log _{e} 2\right)}$ if $\mu=1$ |
| (D) | $P(\widehat{M} \leq 4)=1-e^{-n\left(2 \log _{e} 2-1\right)}$ if $\mu=1$ |
|  |  |


| Q.40 | Let $X_{1}, X_{2}, \ldots, X_{10}$ be a random sample of size 10 from a population having <br> $N\left(0, \theta^{2}\right)$ distribution, where $\theta>0$ is an unknown parameter. <br> Let $T=\frac{1}{10} \sum_{i=1}^{10} X_{i}^{2}$. If the mean square error of $c T(c>0)$, as an estimator of <br> $\theta^{2}$, is minimized at $c=c_{0}$, then the value of $c_{0}$ equals |
| :--- | :--- |
| (A) | $\frac{5}{6}$ |
| (B) | $\frac{2}{3}$ |
| (C) | $\frac{3}{5}$ |
|  | $\frac{1}{2}$ |
|  |  |


| Q.41 | Suppose that $\boldsymbol{X}_{1}, \boldsymbol{X}_{2}, \ldots, \boldsymbol{X}_{10}$ are independent and identically distributed random <br> vectors each having $N_{2}(\boldsymbol{\mu}, \Sigma)$ distribution, where $\Sigma$ is non-singular. If <br>  <br>  <br>  <br> where $\overline{\boldsymbol{X}}=\frac{1}{10} \sum_{i=1}^{10} \boldsymbol{X}_{i}$, then the value of $\log _{e} P\left(U \leq \frac{1}{2}\right)$ equals <br> (A) <br> (B) <br> (C) <br> -10 <br> (D) <br> -1 |
| :--- | :--- |


| Q. 42 | Suppose that $(X, Y)$ has joint probability mass function $\begin{gathered} P(X=0, Y=0)=P(X=1, Y=1)=\theta \\ P(X=1, Y=0)=P(X=0, Y=1)=\frac{1}{2}-\theta \end{gathered}$ <br> where $0 \leq \theta \leq \frac{1}{2}$ is an unknown parameter. Consider testing $H_{0}: \theta=\frac{1}{4}$ against $H_{1}: \theta=\frac{1}{3}$, based on a random sample $\left\{\left(X_{1}, Y_{1}\right),\left(X_{2}, Y_{2}\right), \ldots,\left(X_{n}, Y_{n}\right)\right\}$ from the above probability mass function. Let $M$ be the cardinality of the set $\left\{i: X_{i}=Y_{i}, 1 \leq i \leq n\right\}$. If $m$ is the observed value of $M$, then which one of the following statements is true? |
| :---: | :---: |
|  |  |
| (A) | The likelihood ratio test rejects $H_{0}$ if $m>c$ for some $c$ |
| (B) | The likelihood ratio test rejects $H_{0}$ if $m<c$ for some $c$ |
| (C) | The likelihood ratio test rejects $H_{0}$ if $c_{1}<m<c_{2}$ for some $c_{1}$ and $c_{2}$ |
| (D) | The likelihood ratio test rejects $H_{0}$ if $m<c_{1}$ or $m>c_{2}$ for some $c_{1}$ and $c_{2}$ |
|  |  |



| Q.45 | Let $f$ be a continuous function from [0, 1] to the set of all real numbers. Then <br> which one of the following statements is NOT true? |
| :--- | :--- |
| (A) | For any sequence $\left\{x_{n}\right\}_{n \geq 1}$ in $[0,1], \sum_{n=1}^{\infty} \frac{f\left(x_{n}\right)}{n^{2}}$ is absolutely convergent |
| (B) | If $\|f(x)\|=1$ for all $x \in[0,1]$, then $\left\|\int_{0}^{1} f(x) d x\right\|=1$ |
| (C) | If $\left\{x_{n}\right\}_{n \geq 1}$ is a sequence in $[0,1]$ such that $\left\{f\left(x_{n}\right)\right\}_{n \geq 1}$ is convergent, then <br> $\left\{x_{n}\right\}_{n \geq 1}$ is convergent |
| (D)If $f$ is also monotonically increasing, then the image of $f$ is given by <br> $[f(0), f(1)]$ |  |
|  |  |


| Q.46 | Let $X$ be a random variable with cumulative distribution function |
| :--- | :--- |
|  | $F(x)= \begin{cases}0 & \text { if } x<-1 \\ \frac{1}{4}(x+1) & \text { if }-1 \leq x<0 \\ \frac{1}{4}(x+3) & \text { if } 0 \leq x<1 \\ 1 & \text { if } x \geq 1 .\end{cases}$ |
| (A) | $\lim _{n \rightarrow \infty} P\left(-\frac{1}{2}+\frac{1}{n}<X<\frac{1}{n}\right)=\frac{5}{8}$ |
| (B) | $\lim _{n \rightarrow \infty} P\left(-\frac{1}{2}-\frac{1}{n}<X<\frac{1}{n}\right)=\frac{5}{8}$ |
| (C) | $\lim _{n \rightarrow \infty} P\left(X=\frac{1}{n}\right)=\frac{1}{2}$ |
|  |  |
|  |  |
|  |  |


| Q.47 | Let $(X, Y)$ have joint probability mass function |
| :--- | :--- |
|  | $p(x, y)= \begin{cases}\frac{c}{2^{x+y+2}} & \text { if } x=0,1,2, \ldots ; \quad y=0,1,2, \ldots ; \quad x \neq y \\ 0 & \text { otherwise. }\end{cases}$ |
| Then which one of the following statements is true? |  |
| (A) | $c=\frac{1}{2}$ |
| (B) | $c=\frac{1}{4}$ |
| (C) | $X>1$ |
| and $Y$ are independent |  |
|  |  |


| Q. 48 | Let $\boldsymbol{X}_{1}, \boldsymbol{X}_{2}, \ldots, \boldsymbol{X}_{10}$ be a random sample of size 10 from a $N_{3}(\boldsymbol{\mu}, \Sigma)$ distribution, where $\boldsymbol{\mu}$ and non-singular $\Sigma$ are unknown parameters. If $\begin{array}{cc} \overline{\boldsymbol{X}}_{1}=\frac{1}{5} \sum_{i=1}^{5} \boldsymbol{X}_{i}, & \overline{\boldsymbol{X}}_{2}=\frac{1}{5} \sum_{i=6}^{10} \boldsymbol{X}_{i} \\ S_{1}=\frac{1}{4} \sum_{i=1}^{5}\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}_{1}\right)\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}_{1}\right)^{T}, & S_{2}=\frac{1}{4} \sum_{i=6}^{10}\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}_{2}\right)\left(\boldsymbol{X}_{i}-\overline{\boldsymbol{X}}_{2}\right)^{T}, \end{array}$ <br> then which one of the following statements is NOT true? |
| :---: | :---: |
|  |  |
| (A) | $\frac{5}{6}\left(\overline{\boldsymbol{X}}_{1}-\boldsymbol{\mu}\right)^{T} S_{1}^{-1}\left(\overline{\boldsymbol{X}}_{1}-\boldsymbol{\mu}\right)$ follows a $F$-distribution with 3 and 2 degrees of freedom |
| (B) | $\frac{6}{5\left(\overline{\boldsymbol{X}}_{1}-\boldsymbol{\mu}\right)^{T} S_{1}^{-1}\left(\overline{\boldsymbol{X}}_{1}-\boldsymbol{\mu}\right)} \quad$ follows a $F$-distribution with 2 and 3 degrees of freedom |
| (C) | $4\left(S_{1}+S_{2}\right)$ follows a Wishart distribution of order 3 with 8 degrees of freedom |
| (D) | $5\left(S_{1}+S_{2}\right)$ follows a Wishart distribution of order 3 with 10 degrees of freedom |
|  |  |


| Q. 49 | Which of the following sets is/are countable? |
| :---: | :---: |
| (A) | The set of all functions from $\{1,2,3, \ldots, 10\}$ to the set of all rational numbers |
| (B) | The set of all functions from the set of all natural numbers to $\{0,1\}$ |
| (C) | The set of all integer valued sequences with only finitely many non-zero terms |
| (D) | The set of all integer valued sequences converging to 1 |
| Q. 50 | For a given real number $a$, let $a^{+}=\max \{a, 0\}$ and $a^{-}=\max \{-a, 0\}$. If $\left\{x_{n}\right\}_{n \geq 1}$ is a sequence of real numbers, then which of the following statements is/are true? |
| (A) | If $\left\{x_{n}\right\}_{n \geq 1}$ converges, then both $\left\{x_{n}^{+}\right\}_{n \geq 1}$ and $\left\{x_{n}^{-}\right\}_{n \geq 1}$ converge |
| (B) | If $\left\{x_{n}\right\}_{n \geq 1}$ converges to 0 , then both $\left\{x_{n}^{+}\right\}_{n \geq 1}$ and $\left\{x_{n}^{-}\right\}_{n \geq 1}$ converge to 0 |
| (C) | If both $\left\{x_{n}^{+}\right\}_{n \geq 1}$ and $\left\{x_{n}^{-}\right\}_{n \geq 1}$ converge, then $\left\{x_{n}\right\}_{n \geq 1}$ converges |
| (D) | If $\left\{x_{n}^{2}\right\}_{n \geq 1}$ converges, then both $\left\{x_{n}^{+}\right\}_{n \geq 1}$ and $\left\{x_{n}^{-}\right\}_{n \geq 1}$ converge |
|  |  |


| Q.51 | Let $A$ be a $3 \times 3$ real matrix such that $A\left[\begin{array}{l}0 \\ 1 \\ 1\end{array}\right]=\left[\begin{array}{l}4 \\ 0 \\ 0\end{array}\right], A\left[\begin{array}{l}1 \\ 0 \\ 1\end{array}\right]=\left[\begin{array}{l}0 \\ 1 \\ 1 \\ 0\end{array}\right]=\left[\begin{array}{l}0 \\ 0 \\ 4\end{array}\right]$. Then which of the following statements is/are true? |
| :--- | :--- |
| (A) | $A\left[\begin{array}{l}1 \\ 0 \\ 0\end{array}\right]=\left[\begin{array}{l}2 \\ 2 \\ -2\end{array}\right]$ |
| (B) | $A\left[\begin{array}{l}0 \\ 1 \\ 0\end{array}\right]=\left[\begin{array}{c}2 \\ -2 \\ 2\end{array}\right]$ |
| (C) | $A\left[\begin{array}{l}1 \\ 1 \\ 1\end{array}\right]=\left[\begin{array}{l}2 \\ 0 \\ 2\end{array}\right]$ |
| (D) | $A\left[\begin{array}{l}1 \\ 2 \\ 3\end{array}\right]=\left[\begin{array}{l}8 \\ 4 \\ 0\end{array}\right]$ |


| Q. 52 | Let $X$ be a positive valued continuous random variable with finite mean. If $Y=[X]$, the largest integer less than or equal to $X$, then which of the following statements is/are true? |
| :---: | :---: |
| (A) | $P(Y \leq u) \leq P(X \leq u)$ for all $u \geq 0$ |
| (B) | $P(Y \geq u) \leq P(X \geq u)$ for all $u \geq 0$ |
| (C) | $E(X)<E(Y)$ |
| (D) | $E(X)>E(Y)$ |
| Q. 53 | Let $X$ be a random variable with probability density function $f(x)=\left\{\begin{array}{cl} e^{-x} & \text { if } x \geq 0 \\ 0 & \text { otherwise } \end{array}\right.$ <br> For $a<b$, if $U(a, b)$ denotes the uniform distribution over the interval $(a, b)$ then which of the following statements is/are true? |
| (A) | $e^{-X}$ follows $U(-1,0)$ distribution |
| (B) | $1-e^{-X}$ follows $U(0,2)$ distribution |
| (C) | $2 e^{-X}-1$ follows $U(-1,1)$ distribution |
| (D) | The probability mass function of $Y=[X]$ is $P(Y=k)=\left(1-e^{-1}\right) e^{-k} \quad \text { for } \quad k=0,1,2, \ldots$ <br> where $[x]$ denotes the largest integer not exceeding $x$ |


| Q. 54 | Suppose that $X$ is a discrete random variable with the following probability mass <br> function |
| :--- | :--- |
| $\qquad$ $P(X=0)=\frac{1}{2}\left(1+e^{-1}\right)$ <br>  Which of the following statements is/are true? |  |
| (A) | $E(X)=\frac{e^{-1}}{2 k!}$ for $k=1,2,3, \ldots$. |
| (B) | $E(X)<1$ |
| (C) | $E(X \mid X>0)<\frac{1}{2}$ |
| (D) | $E(X \mid X>0)>\frac{1}{2}$ |
|  |  |


| Q. 55 | Suppose that $U$ and $V$ are two independent and identically distributed random variables each having probability density function $f(x)= \begin{cases}\lambda^{2} x e^{-\lambda x} & \text { if } x>0 \\ 0 & \text { otherwise }\end{cases}$ <br> where $\lambda>0$. Which of the following statements is/are true? |
| :---: | :---: |
|  |  |
| (A) | The distribution of $U-V$ is symmetric about 0 |
| (B) | The distribution of $U V$ does not depend on $\lambda$ |
| (C) | The distribution of $\frac{U}{V}$ does not depend on $\lambda$ |
| (D) | The distribution of $\frac{U}{V}$ is symmetric about 1 |
|  |  |


| Q. 56 | Let $(X, Y)$ have joint probability mass function $p(x, y)=\left\{\begin{array}{cl} \frac{e^{-2}}{x!(y-x)!} & \text { if } x=0,1,2, \ldots, y ; \quad y=0,1,2, \ldots \\ 0 & \text { otherwise } . \end{array}\right.$ <br> Then which of the following statements is/are true? |
| :---: | :---: |
|  |  |
| (A) | $E(X \mid Y=4)=2$ |
| (B) | The moment generating function of $Y$ is $e^{2\left(e^{v}-1\right)}$ for all $v \in \mathbb{R}$ |
| (C) | $E(X)=2$ |
| (D) | The joint moment generating function of $(X, Y)$ is $e^{-2+\left(1+e^{u}\right) e^{v}}$ for all $(u, v) \in \mathbb{R}^{2}$ |
|  |  |


| Q. 57 | Let $\left\{X_{n}\right\}_{n \geq 1}$ be a sequence of independent and identically distributed random variables with mean 0 and variance 1 , all of them defined on the same probability space. For $n=1,2,3, \ldots$, let $Y_{n}=\frac{1}{n}\left(X_{1} X_{2}+X_{3} X_{4}+\cdots+X_{2 n-1} X_{2 n}\right)$ <br> Then which of the following statements is/are true? |
| :---: | :---: |
|  |  |
| (A) | $\left\{\sqrt{n} Y_{n}\right\}_{n \geq 1}$ converges in distribution to a standard normal random variable |
| (B) | $\left\{Y_{n}\right\}_{n \geq 1}$ converges in $2^{\text {nd }}$ mean to 0 |
| (C) | $\left\{Y_{n}+\frac{1}{n}\right\}_{n \geq 1}$ converges in probability to 0 |
| (D) | $\left\{X_{n}\right\}_{n \geq 1}$ converges almost surely to 0 |
| Q. 58 | Consider the following regression model $y_{t}=\alpha_{0}+\alpha_{1} t+\alpha_{2} t^{2}+\epsilon_{t}, \quad t=1,2, \ldots, 100$ <br> where $\alpha_{0}, \alpha_{1}$ and $\alpha_{2}$ are unknown parameters and $\epsilon_{t}$ 's are independent and identically distributed random variables each having $N(\mu, 1)$ distribution with $\mu \in \mathbb{R}$ unknown. Then which of the following statements is/are true? |
| (A) | There exists an unbiased estimator of $\alpha_{1}$ |
| (B) | There exists an unbiased estimator of $\alpha_{2}$ |
| (C) | There exists an unbiased estimator of $\alpha_{0}$ |
| (D) | There exists an unbiased estimator of $\mu$ |


| Q. 59 | Consider the orthonormal set $\left\{v_{1}=\left[\begin{array}{c} \frac{1}{\sqrt{3}} \\ -\frac{1}{\sqrt{3}} \\ \frac{1}{\sqrt{3}} \end{array}\right], v_{2}=\left[\begin{array}{c} \frac{1}{\sqrt{6}} \\ \frac{2}{\sqrt{6}} \\ \frac{1}{\sqrt{6}} \end{array}\right], v_{3}=\left[\begin{array}{c} \frac{1}{\sqrt{2}} \\ 0 \\ -\frac{1}{\sqrt{2}} \end{array}\right]\right\}$ <br> with respect to the standard inner product on $\mathbb{R}^{3}$. If $u=\left[\begin{array}{l}a \\ b \\ c\end{array}\right]$ is the vector such that inner products of $u$ with $v_{1}, v_{2}$ and $v_{3}$ are 1,2 and 3 , respectively, then $a^{2}+b^{2}+c^{2}$ (in integer) equals $\qquad$ |
| :---: | :---: |
|  |  |
|  |  |
| Q. 60 | Consider the probability space $(\Omega, \mathcal{G}, P)$, where $\Omega=\{1,2,3,4\}$, $\mathcal{G}=\{\phi, \Omega,\{1\},\{4\},\{2,3\},\{1,4\},\{1,2,3\},\{2,3,4\}\}$, and $P(\{1\})=\frac{1}{4}$. <br> Let $X$ be the random variable defined on the above probability space as $X(1)=1, X(2)=X(3)=2$ and $X(4)=3$. If $P(X \leq 2)=\frac{3}{4}$, then $P(\{1,4\})$ (rounded off to two decimal places) equals $\qquad$ |
| Q. 61 | Let $\left\{X_{n}\right\}_{n \geq 1}$ be a sequence of independent and identically distributed random variables each having probability density function $f(x)= \begin{cases}e^{-x} & \text { if } x>0 \\ 0 & \text { otherwise }\end{cases}$ <br> For $n \geq 1$, let $Y_{n}=\left\|X_{2 n}-X_{2 n-1}\right\|$. If $\bar{Y}_{n}=\frac{1}{n} \sum_{i=1}^{n} Y_{i}$ for $n \geq 1$ and $\left\{\sqrt{n}\left(e^{-\bar{Y}_{n}}-e^{-1}\right)\right\}_{n \geq 1}$ converges in distribution to a normal random variable with mean 0 and variance $\sigma^{2}$, then $\sigma^{2}$ (rounded off to two decimal places) equals $\qquad$ |
|  |  |



| Q. 65 | Let $\{0.13,0.12,0.78,0.51\}$ be a realization of a random sample of size 4 from a population with cumulative distribution function $F(\cdot)$. Consider testing $H_{0}: F=F_{0} \quad \text { against } \quad H_{1}: F \neq F_{0}$ <br> where $F_{0}(x)= \begin{cases}0 & \text { if } x<0 \\ x & \text { if } 0 \leq x<1 \\ 1 & \text { if } x \geq 1\end{cases}$ <br> Let $D$ denote the Kolmogorov-Smirnov test statistic. If $P(D>0.669)=0.01$ under $H_{0}$ and $\psi= \begin{cases}1 & \text { if } H_{0} \text { is accepted at level } 0.01 \\ 0 & \text { otherwise }\end{cases}$ <br> then based on the given data, the observed value of $D+\psi$ (rounded off to two decimal places) equals $\qquad$ |
| :---: | :---: |

## END OF QUESTION PAPER

